# LONG PATHS IN SPARSE RANDOM GRAPHS

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## Dedicated to Tibor Gallai on his seventieth birthday

Received 8 April 1982

We consider random graphs with n labelled vertices in which edges are chosen independently and with probability c/n. We prove that almost every random graph of this kind contains a path of length  $\ge (1-\alpha(c))n$  where  $\alpha(c)$  is an exponentially decreasing function of c.

Given  $n \in \mathbb{N}$  and  $0 , the probability space <math>\mathcal{G}(n, P(\text{edge}) = p)$  consists of all graphs with a fixed set of n labelled vertices, in which the edges are chosen independently and with probability p. As customary, we say that almost every (a.e.) random graph  $G_{n,p}$  has a property Q if the probability that a graph  $G \in \mathcal{G}(n, P(\text{edge}) = p)$  has Q tends to 1 as  $n \to \infty$ .

Some important recent results in the theory of random graphs concern the existence of long paths in a.e. random graph  $G_{n,c/n}$ , where c is a constant. (For basic properties of graphs and random graphs see [2].) In order to formulate these results, for c>0 set

 $1-\alpha(c)=\sup{\{\alpha\in\mathbf{R}:\ a.e.\ G_{n,\,c/n}\ \text{contains a path of length at least }\alpha n\}},$ 

 $1-\beta(c) = \sup \{\beta \in \mathbb{R}: \text{ a.e. } G_{n,c/n} \text{ contains a cycle of length at least } \beta n \}.$ 

Clearly  $0 \le \alpha(c) \le \beta(c) \le 1$  for every c, and  $\alpha(c)$ ,  $\beta(c)$  are decreasing functions of c.

Erdős and Rényi [3] proved that in a.e.  $G_{n,1/n}$  the largest component has  $O(n^{2/3})$  vertices, so  $\alpha(c)=1$  for  $c \le 1$ . Ajtai, Komlós and Szemerédi [1] proved the somewhat surprising result that  $\alpha(c)<1$  for every c>1, and de la Véga [5] showed independently that  $\alpha(c) \le c_0/c$  for every c>0 and some absolute constant  $c_0$ . Our aim is to show that this bound on  $\alpha(c)$  can be replaced by an exponential function of c.

It is easily seen that  $\alpha(c)$  cannot decay faster than an exponential function of c. Indeed, Erdős and Rényi [3] showed that for every c > 1 a.e.  $G_{n,c/n}$  is such that the order of its largest component is asymptotically  $\left\{1 - \frac{1}{c} \sum_{k=1}^{\infty} k^{k-1} (ce^{-c})^k / k!\right\} n$ . Hence

$$\alpha(c) \ge \frac{1}{c} \sum_{k=1}^{\infty} k^{k-1} (ce^{-c})^k / k! \ge e^{-c}.$$

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**Theorem.** For c>2 a.e.  $G_{n,c/n}$  contains a cycle of length at least  $(1-c^{24}e^{-c/2})n$ . In particular,

$$\alpha(c) \leq \beta(c) \leq c^{24}e^{-c/2}.$$

**Proof.** We may assume that c>200 since if  $c^{24} < e^{c/2}$ , there is nothing to prove. Let  $k, n \in \mathbb{N}$ , k < n, and let V be a set of n labelled vertices. Denote by  $\widetilde{\mathcal{G}}_{k-\text{out}}(n)$  the set of all directed graphs with vertex set V in which the outdegree of every vertex is k:  $d^+(x)=k$  for every  $x \in V$ . Turn  $\widetilde{\mathcal{G}}_{k-\text{out}}(n)$  into a probability space by giving all elements the same probability.

Given a directed graph  $\vec{G}$ , denote by  $\theta(\vec{G})$  the graph having the same vertex set as  $\vec{G}$ , in which xy is an edge iff at least one of  $\vec{xy}$  and  $\vec{yx}$  is an edge of  $\vec{G}$ . Set  $\mathscr{G}_{k\text{-out}}(n) = \{\theta(\vec{G}) : \vec{G} \in \mathscr{G}_{k\text{-out}}(n)\}$  and turn  $\mathscr{G}_{k\text{-out}}(n)$  into a probability space by giving  $G \in \mathscr{G}_{k\text{-out}}(n)$  the probability of the set  $\theta^{-1}(G) \subset \mathscr{G}_{k\text{-out}}(n)$ .

Our proof is based on a recent deep theorem of Fenner and Frieze [4] stating that if  $k \ge 23$  then a.e.  $G \in \mathcal{G}_{k-\text{out}}(n)$  is Hamiltonian. In order to make use of this result, we shall represent  $\mathcal{G}(n, P(\text{edge}) = c/n)$  as the image of a space of random directed graphs.

Define r=d/n, 0 < r < 1, by

$$c/n = 2r - r^2.$$

Then clearly

$$c/2 + c^2/8n < d < c/2 + c^2/7n$$

if n is sufficiently large. Let  $\mathcal{G}(n, r)$  be the probability space of directed graphs with a fixed set V of n labelled vertices in which the edges are chosen independently and with probability r. Then in  $\theta(G)$  the edges are chosen independently and with probability

$$1-(1-r)^2=2r-r^2=c/n,$$

so  $\{\theta(\vec{G}): \vec{G} \in \mathcal{G}(n,r)\}$ , with probability inherited from  $\vec{\mathcal{G}}(n,r)$ , is just another representation of  $\mathcal{G}(n, P(\text{edge}) = c/n)$ . Hence it suffices to show that a.e.  $\vec{G} \in \mathcal{G}(n,r)$  is such that  $\theta(\vec{G})$  contains a cycle of length at least  $(1-c^{24}e^{-c/2})n$ .

We shall do this by showing that a.e.  $\vec{G} \in \vec{\mathcal{G}}(n,r)$  is such that every vertex of a large set W dominates at least 23 vertices of W. Then the Fenner—Frieze theorem applied to the random graph  $\vec{G}[W]$  will imply that a.e.  $\vec{G} \in \vec{\mathcal{G}}(n,r)$  is such that  $\theta(\vec{G})$  contains a cycle of length at least W. We shall look for an appropriate set W by omitting from V disjoint sets  $U_0, U_1, \ldots, U_t$ . The set  $U_0$  will consist of all vertices of small outdegree (together with the vertices of too large outdegree, but this is only for the sake of convenience), then  $U_1$  will be the set of vertices of small outdegree in  $G - U_0$ , and so on.

Given  $x \in V$ , the outdegree  $d^+(x)$  of x in a random directed graph  $\vec{G}_{n,r}$  has binomial distribution with parameters n-1 and r=d/n. Let us estimate  $P(d^+(x) \le 24)$  and  $P(d^+(x) \ge 6d)$ . Since c > 104, if n is sufficiently large then

$$(1-d/n)^{n-1-i} < \exp\left\{-d+25 d/n\right\} < \exp\left\{-c/2-c^2/8n+13c/n\right\} < e^{-c/2}$$

for i < 25. Consequently

$$P(d^+(x) \le 24) = \sum_{i=0}^{24} {n-1 \choose i} \left(\frac{d}{n}\right)^i (1 - d/n)^{n-1-i} < e^{-c/2} \sum_{i=0}^{24} (c/2)^i / i!.$$

Also,

$$P(d^+(x) \ge 6d) < \sum_{i=\lceil 6d \rceil}^{n-1} {n-1 \choose i} \left(\frac{d}{n}\right)^i < \sum_{i=\lceil 6d \rceil}^{\infty} d^i/i! < \sum_{i=\lceil 6d \rceil}^{\infty} (e/6)^i < e^{-c}.$$

For  $\vec{G} \in \vec{\mathcal{G}}(n,r)$  let

$$U_0 = U_0(\vec{G}) = \{x \in V : d^+(x) \le 24 \text{ or } d^+(x) \ge 6d\}.$$

Then  $|U_0|$  has binomial distribution with parameters n and p, where p=p(n) satisfies

$$p < e^{-c} + e^{-c/2} \sum_{i=0}^{24} (c/2)^i / i! = p_0.$$

Let us note the following simple property of the binomial distribution  $S_{n,p}$  with parameters n and p: if  $0 and <math>(pn)^{-1/2} < \varepsilon \le 1/2$  then

(1) 
$$P(|S_{n,n}-pn| \ge \varepsilon pn) < e^{-\varepsilon^2 pn/3}.$$

By applying (1) we find that

$$P\left(|U_0| \ge \frac{3}{2} p_0 n\right) < e^{-p_0 n/12} = P_0.$$

Now let us define disjoint sets  $U_1, U_2, ...$ , all functions of a random directed graph  $\vec{G} \in \vec{\mathscr{G}}(n, r)$ , as follows. Suppose  $i \ge 1$  and we have defined  $U_0, U_1, ..., U_{i-1}$ . Set

$$U_i = \left\{ x \in V - \bigcup_{i=0}^{i-1} U_j \colon \left| \Gamma_+(x) \cap \bigcup_{i=0}^{i-1} U_j \right| \ge 2 \right\},\,$$

where

$$\Gamma_+(x) = \{ y \in V \colon \vec{xy} \in \vec{E}(\vec{G}) \}.$$

Let 
$$D = 9d^2$$
,  $t = \left[ \frac{2}{3} \log n / \log (1/Dp_0) \right]$  and

$$p_i = D^j p_0^{j+1}, \quad j = 1, 2, ..., t.$$

It can be checked that the assumptions we have made imply that  $Dp_0 < 1/4$  and so  $p_j < 4^{-j-1}$ , j=0, 1, ..., t. We shall show that the probability that  $|U_j| \ge \frac{3}{2} p_j n$ ,

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conditional on  $|U_i| \le \frac{3}{2} p_i n$ , i = 0, 1, ..., j-1, is rather small. To be precise,

(2) 
$$P\left(|U_j| \ge \frac{3}{2} p_j n \, \middle| \, |U_i| \le \frac{3}{2} p_i n, \quad i = 0, 1, ..., j-1\right) \le e^{-p_j n/12} = P_j,$$

$$j = 1, 2, ..., t.$$

In order to prove (2), consider disjoint sets  $V_0, V_1, ..., V_{j-1}$  of V satisfying

$$v_i = |V_i| \le \frac{3}{2} p_i n, \quad i = 0, 1, ..., j-1.$$

Let  $x_0 \in V - \bigcup_{i=0}^{j-1} V_i$ . We claim that

(3) 
$$P\left(\left|\Gamma_{+}(x_{0})\cap\bigcup_{i=0}^{j-1}V_{i}\right|\geq2\left|U_{i}=V_{i},\quad i=0,\,1,\,...,\,j-1\right)\leq p_{j},\right)$$

j=1, 2, ..., t. Note that, by (1), inequality (3) implies (2), so it suffices to prove (3). First consider j=1. Since  $P(x \in U_0) < 1/2$ ,

$$\begin{split} P(|\Gamma_{+}(x_{0}) \cap V_{0}| &\geq 2 \left| U_{0} = V_{0} \right) \leq 2 \sum_{i=25}^{\lfloor 6d \rfloor} \sum_{k=2}^{i} \binom{v_{0}}{k} \binom{n-1-v_{0}}{i-k} r^{i} (1-r)^{n-1-i} \\ &\leq 3 \sum_{i=25}^{\lfloor 6d \rfloor} \sum_{k=2}^{i} \left( \frac{3}{2} p_{0} \right)^{k} \frac{d^{i}}{k! (i-k)!} e^{-d} \leq 3 e^{-d} \sum_{k=2}^{\lfloor 6d \rfloor} \frac{(3p_{0} d/2)^{k}}{k!} \sum_{i=k}^{\infty} d^{i-k} / (i-k)! \leq 5 d^{2} p_{0}^{2} < p_{1}. \end{split}$$

The last but one inequality above holds because  $p_0 d < 1/5$ .

Now assume that  $2 \le j \le t$  and we have proved (3) and (2) for smaller values of j. If a vertex x belongs to  $U_j$  then x dominates at least one vertex of  $U_{j-1}$  and at least two vertices of  $\bigcup_{i=0}^{j-1} U_i$ . Taking into account that  $\frac{3}{2} \sum_{i=0}^{j-1} p_i < 2p_0 < 1/2$ , we find that for  $x_0 \notin \bigcup_{i=0}^{j-1} V_i$ 

$$\begin{split} &P\left(\left|\Gamma_{+}(x_{0})\bigcap\bigcup_{i=0}^{j-1}U_{i}\right|\geq2\left|U_{i}=V_{i},\quad i=0,1,...,j-1\right)\leq\\ &\leq2\sum_{k=25}^{\lfloor 6d\rfloor}v_{j-1}\sum_{i=0}^{j-1}v_{i}\binom{n-2}{k-2}r^{k}(1-r)^{n-1-k}\leq\\ &\leq3\sum_{k=25}^{\lfloor 6d\rfloor}\left(\frac{3}{2}p_{j-1}\right)(2p_{0})\frac{d^{k}}{(k-2)!}e^{-d}\leq9d^{2}p_{0}p_{j-1}=p_{j}. \end{split}$$

This proves (3) and so (2).

The set  $\bigcup_{i=0}^{r} U_i$  will not always do for the set U = V - W of exceptional vertices.

Set

$$U = \left\{ x \in V - \bigcup_{j=1}^{t} U_j \colon |\Gamma_+(x) \cap U_t| \ge 2 \right\}.$$

Let us estimate the probability of U being large, conditional on  $U_0, U_1, \ldots, U_t$  being not too large. As before, consider disjoint sets  $V_0, V_1, \ldots, V_t$  of V satisfying

$$v_j = |V_j| \le \frac{3}{2} p_j n, \quad j = 0, 1, ..., t.$$

Then for  $x_0 \in V - \bigcup_{i=0}^t V_i$ 

$$P(|\Gamma_{+}(x_0) \cap V_t| \ge 2 |U_i = V_i, \quad i = 0, 1, ..., t) \le$$

$$\leq 2 \sum_{i=25}^{\lfloor 6d \rfloor} {v_t \choose 2} {n-2 \choose i-2} r^i (1-r)^{n-1-i} \leq 2d^2 p_t^2 \leq (Dp_0)^{2(t+1)} \leq n^{-4/3}.$$

Hence

$$(4) P\left(\widetilde{U}\neq\emptyset\big||U_i|\leq\frac{3}{2}p_in, \quad i=0,1,...,t\right)\leq n^{-1/3}.$$

Putting together (2) and (4) we find that

$$P(|U_j| \le \frac{3}{2}p_j n, \quad j = 0, 1, ...t, \text{ and } \widetilde{U} = \emptyset) \ge 1 - \sum_{i=0}^t P_j - n^{-1/3} = 1 - o(1).$$

Now let W be a set of maximum cardinality for which

$$|\Gamma_+(x)\cap W|\geq 23$$

for every  $x \in V - W$ . If  $\widetilde{U} = \emptyset$  then the set  $V - \bigcup_{j=0}^{t} U_j$  will do for W since a vertex not in  $\bigcup_{j=0}^{t} U_j$  dominates at most one vertex in  $U_t$  and at most one vertex in  $\bigcup_{j=0}^{t-1} U_j$ , so it dominates at least 25-2=23 vertices in  $V - \bigcup_{j=0}^{t} U_j$ .

Consequently

$$P(|W| \ge (1-2p_0)n) \ge P(|U_j| \le \frac{3}{2}p_j n, \quad j=0,...,t, \text{ and } \widetilde{U}=\emptyset) = 1-o(1).$$

Having found this large set W the proof is almost complete. Indeed, if  $W_0$  is a fixed set of vertices with  $|W_0| \ge (1-2p_0)n$  then

$$P(\theta(\vec{G})[W])$$
 is Hamiltonian  $|W = W_0| = 1 - o(1)$ .

Consequently a.e.  $\vec{G} \in \mathcal{G}(n, r)$  is such that  $\theta(\vec{G})$  contains a cycle of length at least  $(1-2p_0)n$ .

Note that in the proof we were very generous in our estimate of  $p_0$ . This is because the correct upper bound for  $\beta(c)$  will be  $c_1ce^{-c}$  where  $c_1$  is an absolute constant. We shall return to this in a later note.

### References

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